

Human capital accumulation in temporary jobs: specific or general?

Work in progress

F. Berton, University of Eastern Piedmont and LABOR

F. Devicienti, University of Turin, LABOR and Collegio Carlo Alberto

L. Pacelli, University of Turin and LABOR

Flexibilization

- Flexibilization has been the main labor market policy objective during the last 20 years
- pursued by allowing temporary hires under less and less binding conditions
- Labor market deregulation implies less **job security** (the capability to stay with a certain job and employer)
- And it demands more **employment security** (the capability to stay employed across jobs and employers) through:
 - higher arrival rate of job offers
 - easier entries on the labor market
 - easier transitions across employment relationships
 - shorter unemployment
 - accumulation of actual experience

External careers

- Emphasis has indeed been put on **external careers** across firms
- Limited to the initial period on the labour market or extending over time
 - In this perspective empirical research has been so far mostly devoted to evaluate the **port-of-entry hypothesis** (temporary jobs as stepping stones into open-ended ones with respect to unemployment), with mixed results
- But the POE hypothesis only tells a part of the story of external careers
 - it implicitly assumes the primary relevance of job security
 - it may well occur within firms and not across them

Human capital

- An underexplored issue is that of **human capital accumulation and portability** in temporary jobs
- Its existence is crucial in assessing the viability of an external career and hence the mere possibility of acquiring employment security

Specific or general

- The standard theory of human capital allows to make predictions about the kind of human capital accumulated by workers hired under different contracts:
 - Firms may invest in workers' specific HC;
 - The stricter the commitment in the work relationship, the larger the investment.
 - We thus expect permanent workers to get a larger amount of specific HC.
 - *There is evidence in the literature of permanent workers to be more likely to receive firm-financed training.*
 - We expect workers who enjoy a lower degree of job security to have a higher incentive to invest in their general HC
- Specific: “firm” and “occupation” specific, the second part being portable across firms

Transitions across firms

- If j and k are different occupations
 - $\Pr\{\text{trans}(j,j)\} = f(\text{SHC}^+, \text{GHC}^+)$
 - $\Pr\{\text{trans}(j,k)\} = g(\text{GHC}^+)$
- Hence we expect the relative probability to make a transition across different occupations to be higher for temporary workers than for permanent ones. In symbols:
$$\Pr_{\text{temp}}\{\text{trans}(j,k)\} / \Pr_{\text{temp}}\{\text{trans}(j,j)\} > \Pr_{\text{perm}}\{\text{trans}(j,k)\} / \Pr_{\text{perm}}\{\text{trans}(j,j)\}$$
- The possibility for temporary workers to reach employment security through external careers rests (also) on this hypothesis
- It can be empirically tested, what we aim to do in this contribution

Data

- We use WHIP data:
 - panel of individual work histories from administrative data
 - representative of the private sector of Italy's labor market
 - careers are observed monthly
 - the series covers now the period from 1985 to 2004
- Administrative data represent a convenient choice since careers can be observed in details and without measurement errors

Occupation

- However, occupation is observed only at a very aggregate level (blue/white collars)
- we thus ***approximate it through 2-digit sector***
- In order for sector to be a good approximation of occupation we restrict the analysis to ***blue collars of manufacturing and building sectors***

Sample selection

- In order for these workers to be a very homogeneous group and minimize unobserved heterogeneity we further select workers:
 - **entering the labor market** in the period 1998 - 2002
 - with either a **full-time open-ended contract or a full-time direct-hire temporary** one
 - aged **19-25 at entry**
 - at their **first involuntary unemployment spell** (i.e. lasting more than one month)
 - after a **unique employment relationship**

Descriptive statistics

- We end up with a sample of 1379 obs.:
 - 1126 males and 253 females
 - 975 open-ended contracts and 404 temporary contracts
 - median duration of the (unique) previous work experience: 5 months for open-ended-contract workers and 4 months for temporary ones

Exit U	Same sector	Different sector
Open-ended	25.8%	21.2%
temporary	19.8%	31.9%

The model

- We model the unemployment duration in a competing-risk setting in discrete time (mlogit)
- The **dependent variable** takes four values
 - 0 : no exit and censored spells (unemployment spells lasting more than 18 months are assumed to be censored)
 - 1 : exit to the same sector
 - 2 : exit to a different sector
 - 3 : other exits (free lance ...)

Covariates

- The covariates aim at controlling for individual, firm and labor market conditions at entry and for the total amount of actual experience accumulated before separation:
 - matrix X: gender, age at entry, wage at entry
 - matrix Z: sector, firm size
 - matrix W: local youth unemployment rate, geo. area, entry year
 - experience: duration of the (unique) preceding work relationship
- We thus argue that conditional on X, Z, W and experience a worker is hired (and dismissed) under a temporary or a permanent arrangement randomly

Specification of HC

- In the baseline specification a dummy variable D takes value 1 when the preceding contract was temporary
- By using exits to the same sector as benchmark and relative risk ratios, the hypothesis we are testing holds if

$$\beta_{EXIT=2}(D) > 1$$

- Then we allow the effect to evolve over time:

$$Y_{EXIT=2}(Tk / D) > 1$$

where T_k is a dummy on U duration, $k=1$ to K

Results

- In the baseline specification:

Exit to a different sector	RRR	Robust s.e.	z	P>z
D	1.3794	0.27367	1.62	0.105

- The effect is driven by what happens during the first months of unemployment

	RRR	Robust SE	z	P > z
Open-ended contracts				
2 to 4 months	<i>Coefficient normalized to 1</i>			
5 to 7 months	0.8256	0.22283	-0.71	0.478
8 to 10 months	0.5304	0.14271	-2.36	0.018
11 to 14 months	0.5956	0.17726	-1.74	0.082
15 to 18 months	0.8502	0.27095	-0.51	0.611
Temporary contracts				
2 to 4 months	2.2368	0.80983	2.22	0.026
5 to 7 months	1.3487	0.52185	0.77	0.439
8 to 10 months	0.4831	0.15794	-2.23	0.026
11 to 14 months	0.9499	0.40063	-0.12	0.903
15 to 18 months	0.9992	0.55010	0.00	0.999

Results

Prob > chi2		0	Number of obs	14276		
Pseudo R2	0.0354	Wald chi2(66)	292.25			
Log pseudolikelihood	-3609.3334					
Exit to another sector	RRR	Robust SE	z	P > z 	95% Conf. Interval	
Female	1.0153	0.22913	0.07	0.946	0.6523	1.5801
Age at entry	1.0574	0.04343	1.36	0.174	0.9756	1.1460
Wage at entry (hundreds of €)	0.9839	0.02132	-0.75	0.453	0.9430	1.0266
Building sector	0.4179	0.07993	-4.56	0.000	0.2873	0.6080
Firm size at entry	1.0000	0.00002	-0.73	0.466	0.9999	1.0000
Youth unemployment rate	1.0926	0.04892	1.98	0.048	1.0008	1.1928
Youth unemployment rate^2	0.9984	0.00064	-2.56	0.011	0.9971	0.9996
North East	1.0102	0.27404	0.04	0.970	0.5936	1.7191
Center	0.8439	0.22078	-0.65	0.517	0.5054	1.4093
South and main isles	0.7364	0.34276	-0.66	0.511	0.2958	1.8336
Entry in 1999	1.3562	0.32400	1.28	0.202	0.8491	2.1661
Entry in 2000	0.9186	0.23197	-0.34	0.737	0.5600	1.5069
Entry in 2001	0.8925	0.25353	-0.40	0.689	0.5115	1.5574
Entry in 2002	1.0756	0.32661	0.24	0.810	0.5932	1.9504
Actual experience	0.8523	0.05416	-2.51	0.012	0.7525	0.9653
Actual experience^2	1.0121	0.00447	2.73	0.006	1.0034	1.0209
Actual experience^3	0.9998	0.00009	-2.71	0.007	0.9996	0.9999
5 to 7 months	0.7510	0.17018	-1.26	0.206	0.4817	1.1709
8 to 10 months	0.4153	0.09023	-4.04	0.000	0.2712	0.6357
11 to 14 months	0.5423	0.13398	-2.48	0.013	0.3341	0.8801
15 to 18 months	0.6574	0.17959	-1.54	0.125	0.3849	1.1230
Dummy for temporary	1.3794	0.27367	1.62	0.105	0.9350	2.0350

Robustness check

- The effect disappears in sectors and for occupations with more horizontal skills (white collars in services)

	RRR	Robust SE	z	P > z
Open-ended contracts				
2 to 4 months				<i>Coefficient normalized to or</i>
5 to 7 months	0.6313	0.31578	-0.92	0.358
8 to 10 months	0.5275	0.27536	-1.23	0.220
11 to 14 months	0.5092	0.27927	-1.23	0.218
15 to 18 months	0.8018	0.67284	-0.26	0.792
Temporary contracts				
2 to 4 months	0.7728	0.36779	-0.54	0.588
5 to 7 months	0.5736	0.28005	-1.14	0.255
8 to 10 months	0.2756	0.13743	-2.58	0.010
11 to 14 months	0.5843	0.32868	-0.96	0.339
15 to 18 months	2.1983	2.25889	0.77	0.443

Concluding remarks

- Consistently with the theory, we found evidence that temporary jobs entail a higher relative amount of general HC with respect to open-ended jobs
- This should indeed make temporary workers relatively more apt to external careers
- However, there is still a matter of absolute level of human capital, with respect to which temporary workers suffer from a negative gap compared to permanent ones [Bassanini 2007] and that makes transitions (even across employers) easier for workers with open-ended contracts [Berton et al. 2009]

Next steps

We plan to re-estimate our models using additional data sources in which:

- (i) the occupation is observed at a very fine level;
- (ii) workers laid-off by the same firm can be identified, allowing for firm-level effects in estimation;
- (iii) it is possible to observe firms' closure events, as well as to define industry-wide demand shocks (on sales or value added), and therefore to better identify involuntary unemployment.