# Coherent small area estimates for skewed business data EESW 2013

Session 2, Advanced methods for processing and analysis Chair: Maria Dolores Ugarte

Thomas Zimmermann Ralf Münnich

Nuremberg, 9<sup>th</sup> September 2013

# SAE and Business Data

- Small area methods are now in wide use
  - Geographical areas of interest
  - Domains of interest, e.g. NACE classes
- ► Business data characterized by outliers and skewed distributions → violation of assumptions
- Relationships between variables may be multiplicative
- Applying transformations may help to recover some of these assumptions
- Business surveys often based on designs with highly different weights
- Interaction between designs and models is of crucial importance
- Estimates for small areas should be coherent with estimates for aggregates

### Estimators based on transformations

We may assume the following unit-level lognormal-mixed model (Berg and Chandra, 2012)

$$\log(y_{dj}) = \mathbf{x}_{dj}^{\mathsf{T}} \boldsymbol{\beta} + u_d + \varepsilon_{dj}, \quad d = 1, \dots, D, \ j = 1, \dots, N_d$$

where  $\mathbf{x}_{dj}$  includes an intercept and the other components of it are appropriately transformed.  $u_d \stackrel{i.i.d.}{\sim} N(0, \sigma_u^2)$  is the domain-specific random effect and  $\varepsilon_{dj} \stackrel{i.i.d.}{\sim} N(0, \sigma_{\varepsilon}^2)$  the individual error term. The domain-specific random effect is assumed to be independent from the error term.

# An optimal predictor

Minimizing the MSE under the unit-level lognormal mixed model yields the Empirical Bayes predictor

$$\hat{\theta}_{d}^{EBLOG} = \frac{1}{N_{d}} \left( \sum_{j \in s_{d}} y_{dj} + \sum_{j \notin s_{d}} \hat{y}_{dj}^{EBLOG} \right)$$
(1)

derived by Berg and Chandra (2012). The predictions for the non-sampled values ( $j \notin s_d$ ) are given by :

$$\hat{y}_{dj}^{EBLOG} = \exp\left(\mathbf{x}_{dj}^{T}\hat{\boldsymbol{\beta}} + \hat{u}_{d} + 0.5\hat{\sigma}_{\varepsilon}^{2}(\hat{\gamma}_{d}/n_{d} + 1)\right)$$
(2)

with  $\hat{\gamma}_d = rac{\hat{\sigma}_u^2}{\hat{\sigma}_u^2 + \hat{\sigma}_\varepsilon^2/n_d}.$ 

# Area Level Lognormal Model

Assuming that the direct means are lognormally distributed, Slud and Maiti (2006) propose the following predictor:

$$\hat{\theta}_{d}^{ALLOG} = \exp\left(\bar{\mathbf{X}}_{d}^{T}\hat{\boldsymbol{\beta}} + \hat{\boldsymbol{u}}_{d} + 0.5\hat{\sigma}_{u}^{2}\left(1 - \hat{\gamma}_{d}\right)\right)$$
(3)

Estimator (3) corrects for the presence of the random effect but ignores the variability of the parameter estimates.

# Other Estimators

#### Design-based / Model-assisted Estimators

- Direct estimator, which is a weighted sample mean
- Generalized Regression Estimators:

$$\hat{ heta}_{d}^{GREG} = rac{1}{\widehat{N}_{d}} \left[ \sum_{k \in U_{d}} \hat{y}_{k} + \sum_{k \in s_{d}} w_{k} \left( y_{k} - \hat{y}_{k} 
ight) 
ight]$$

GREG Linear fixed-effects model used to predict  $\hat{y}_k$ MLogGREG Predictions  $\hat{y}_k^{EBLOG}$  are used

#### **Benchmarked Estimators**

We benchmark estimator (1) against the weighted sample total for the population to obtain the LOGBench predictor.

# Dataset

- Our dataset is based on
  - the Italian register of enterprises (ASIA 2003)
  - and the survey of small and medium enterprises (PMI)
- ► We focus on the subset of small and medium enterprises → about 4.3 million entries
- Our variable of interest is the mean of labour costs in each domain
- Auxiliary information: Number of employees of each enterprise
- The original datasets were kindly provided by ISTAT

# Setup

- Strata are cross-classifications of the first digit of the industry classification, Italian NUTS 1 areas and the classified size variable in terms of numbers of employees
- As most enterprises in the data set have less than 5 employees, we aggregate the size variable
   Group 1 All enterprises with 1 5 employees
   Group 2 Enterprises with 6 99 employees
- Stratum sizes vary between 799 and 364294
- Focus on SME: no take-all stratum
- Total sample size of n = 67,989
- R = 10,000 simulation runs

#### Domains

We consider two types of domains

### 1. Planned domain structures

Domains as cross-classifications of NUTS 1 and the first digit of the industry classification

D = 45 domains

Domain sizes vary between 6340 and 398874

### 2. Unplanned domain structures

Domains as cross-classifications of Italy's 20 regions and the first digit of the industry classification D = 180 domains

Domain sizes range from 144 to 229873

Setup and Design Results

#### Lehrstuhl für Wirtschafts- und Sozialstatistik

# Gelman Factors and distribution of weights

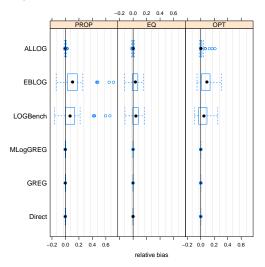
Following Münnich and Burgard (2012) the **Gelman factor** is defined as the ratio of the largest to the smallest (design) weight:

$$\mathsf{GF} = \frac{\max_{i=1,\dots,N} \frac{1}{\pi_i}}{\min_{i=1,\dots,N} \frac{1}{\pi_i}}$$

Allocation	max/min	q95/q05	q75/q25
PROP	1.06	1.01	1.00
EQ	455.33	134.95	6.95
OPT	73.38	41.99	18.55

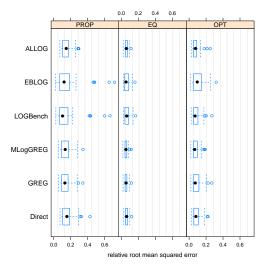
Setup and Design Results

# Relative Bias - planned domains



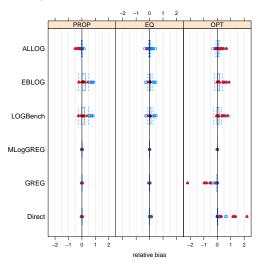
Setup and Design Results

## **RRMSE** - planned domains



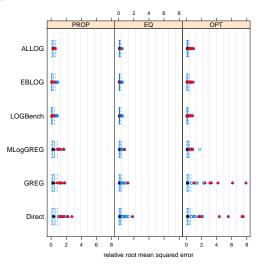
Setup and Design Results

## Relative Bias - unplanned domains



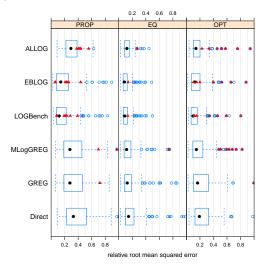
Setup and Design Results

# **RRMSE** - unplanned domains



Setup and Design Results

# **RRMSE** - unplanned domains



# Summary and Outlook

- Model-assisted estimators best choice for planned (large) domain structurs
- For unplanned domain structures model-based estimators help to produce more reliable estimates
- In this application benchmarking is desirable for the estimation at domain level as well
- Incorporating design information may be beneficial for model-based estimators
- MSE estimation for log-transformed estimators is very computerintense

#### Acknowledgement

BLUE-ETS (Enterprise and Trade Statistics) is funded by the European Commission within the Seventh Framework Programme