



# The British Low-Wage Sector and the Employment Prospects of the Unemployed

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- Several studies confirm these concerns, e.g. Stewart & Swaffield 1999, Stewart 2007, Cappellari & Jenkins 2008, Clark & Kanellopoulos 2013.





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**Question:** Does this negative picture of low-wages also hold for the subsample of initially unemployed?







### What is new:

• Analyzing a subsample which contains initially unemployed.



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- Medium-term time frame (up to six years after becoming unemployed).

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- Analyzing a subsample which contains initially unemployed.
- Medium-term time frame (up to six years after becoming unemployed).
- Differentiate the effect of a low-paid job according to individual and job characteristics (Knabe & Plum 2013).
- In the econometric model it is explicitly respected for correlated random effects between the three labor market states (high-paid, low-paid, unemployed).





## The theoretical perspective:

No clear answer:



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#### No clear answer:

• Positive effect on the level of human capital.



## The theoretical perspective:

#### No clear answer:

- Positive effect on the level of human capital.
- Layard, Nickell & Jackman (1991, p. 249): 'While unemployment is a bad signal, being in a low-quality job may well be a worse one'.





# **Existing studies:**





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### **Existing studies:**

- Evidence for low-pay persistence is found (Stewart & Swaffield 1999, Stewart 2007, Clark & Kanellopoulos 2013).
- 'Negative duration dependence' in unemployment (e.g. Kroft, Lange & Notowidigdo 2013).
- Hence: 'the prospects of becoming high-paid might darken when working in the low wage sector but may even be worse when staying unemployed'.









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- Restricted to men.
- Dropping: self-employed, disabled and men attending school or working in the army.
- Age frame: 20-60 years.

# Labor market position



### Labor market position

 ILO definition to differentiate between unemployed and inactive.



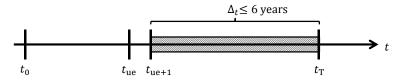


### Labor market position

- ILO definition to differentiate between unemployed and inactive.
- OECD (1997): two third of the median gross hourly wage of both sexes (including paid overtime) as low-pay threshold (annual adjusted)

### Identification of the labor market dynamics

Figure: Identification of labor market dynamics



*Note:*  $t_0$  = first time observed in the sample;  $t_{\rm em}$  = being employed;  $t_{\rm ue}$  = being observed for the first time unemployed after being employed;  $t_{\rm T}$  = up to six years after  $t_{\rm ue}$ . The shadowed box indicates the analyzed time frame.



**Table :** Transition into High-Paid Employment

	First time being high-paid employed					
$\Delta_t$ after	(low-paid in at least one period before)					
${\sf unemployment}^1$	Total	Less than	At least			
		a college degree				
1	95 (-)	59 (-)	36 (-)			
2	34 (17)	21 (12)	13 (5)			
3	14 (12)	12 (11)	2 (1)			
4	12 (11)	8 (7)	4 (4)			
5	2 (2)	- (-)	2 (2)			
6	1 (1)	1 (1)	-(-)			
$\sum$	158 (43)	101 (31)	57 (12)			
Total	210	143	67			
Share	75.23% (20.47%)	78.32% (21.67%)	85.07% (17.91%)			

Source: BHPS waves 8-18, N=796.  $^1$   $\Delta_t$  after unemployment refers to the length  $\Delta_t$ , measured in years, when the initially unemployed man obtains for the first time a high-paid job. Note that the labor market position is observed at one time point in the respective year.

#### **Table:** Control variables

Variables	Description			
Young	Dummy: 1 if observation is 30 years or younger, 0 otherwise			
Old	Dummy: 1 if observation is older than 54 years, 0 otherwise			
Married	Dummy: 1 if observation is married, 0 otherwise			
Health	Dummy: 1 if self reported health status is excellent or good, 0 else			
Unemployment rate	State-level unemployment rate; annual averages; in percent			
Interaction with labor market position				
College-educated	Dummy: 1 if observation obtained a college degree			
	(ISCED 5 or 6), 0 otherwise <sup>1</sup>			
Low social status	Dummy: 1 if presents' job RGSC-value is 4 or above, 0 otherwise <sup>2</sup>			
Robustness <sup>3</sup>				
Female	Dummy: 1 if woman, 0 otherwise			

<sup>&</sup>lt;sup>1</sup> ISCED: International Standard Classification of Education. <sup>2</sup> RGSC: Registrar General's Social Classes

is 1=Professional occ., 2=Managerial & technical occ., 3=Skilled non-manual, 4=Skilled manual,

<sup>5=</sup>Partly skilled occ., 6=Unskilled occ. <sup>3</sup> Only included in the robustness checks.

**Table :** Descriptive Statistics<sup>1</sup>

	Full Sample <sub>t</sub>	$high ext{-}paid_t$	$low-paid_t$	$unemployed_t$
Young	0.274	0.199	0.378	0.340
Old	0.104	0.109	0.098	0.100
Married	0.665	0.744	0.594	0.490
Health	0.687	0.708	0.665	0.650
Unemployment-rate	5.198	5.190	5.247	5.111
College-educated	0.323	0.373	0.240	0.310
Low social class	$0.539^{2}$	0.457	0.681	_
Observations	796	442	254	100

*Source*: BHPS waves 8-18, N = 796. <sup>1</sup> share of observations in the respective group, <sup>2</sup> only including high-paid and low-paid in the full sample.





# **Preliminary remarks**





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- Assuming a first order Markov process.
- Unobserved heterogeneity (Heckman 1981a)
- Initial conditions problem (Heckman 1981b)

#### The two binary outcome variables are defined as:

$$y_{1it} = \begin{cases} 1 & \text{if the person is employed in a high-paid job,} \\ 0 & \text{otherwise,} \end{cases}$$

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$$\begin{array}{rcl} y_{1it} & = & \begin{cases} 1 & \text{if the person is employed in a high-paid job,} \\ 0 & \text{otherwise,} \end{cases} \\ \text{and if } y_{1it} = 0, y_{2it} & = & \begin{cases} 1 & \text{if the person is unemployed,} \\ 0 & \text{otherwise,} \end{cases} \end{array}$$

The observed binary outcome variable is defined as:

$$\begin{array}{lll} y_{1it} & = & \mathbf{1}(x'_{1it}\beta_1 + \gamma_{11}y_{1i(t-1)} + \gamma_{13}y_{3i(t-1)} + \pi_{11}y_{1i0} + \pi_{13}y_{3i0} \\ & & + \overline{x}'_{1i}\delta_1 + \kappa_{1i} + \epsilon_{1it} > 0), \end{array}$$

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and if 
$$y_{1it} = 0$$
,  $y_{2it} = \mathbf{1}(x'_{2it}\beta_2 + \gamma_{21}y_{1i(t-1)} + \gamma_{23}y_{3i(t-1)} + \pi_{21}y_{1i0} + \pi_{23}y_{3i0} + \overline{x}'_{2i}\delta_2 + \kappa_{2i} + \epsilon_{2it} > 0)$ .

### **Correlation structure:**

$$corr(\nu_{1it}, \nu_{1is}) = \begin{cases} \sigma_{\kappa_1}^2 & \text{if } t \neq s, \\ \sigma_{\kappa_1}^2 + 1 & \text{if } t = s, \end{cases}$$

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 Applying a correlated simulated multivariate random effects (CSM RE) probit model.



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- Main feature: the complete variance-covariance matrix is estimated at once (Cappellarie & Jenkins 2006).
- Multivariate normal probability functions of orders higher than two must be simulated.
- For simulation, Halton draws are applied (Train 2003) due to high accuracy and stability (Plum 2013).





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#### Regression results

	RE Probit		CSM RE Probit			
	coeff.	std. err.	coeff.	std. err.		
dependent variable:	employed in a high-paid job in t					
$high ext{-}paid_{t-1}$	0.929	0.237	0.986	0.240		
$\times$ college-educated $t-1$	0.428	0.255	0.434	0.258		
$\times$ low social class <sub>t-1</sub>	-0.201	0.220	-0.182	0.224		
$low-paid_{t-1}$	0.369	0.275	0.566	0.289		
$\times$ college-educated $_{t-1}$	-0.038	0.308	0.037	0.313		
$\times$ low social class <sub>t-1</sub>	-0.257	0.262	-0.228	0.264		
$unemployed_{t-1}$	reference category					
$\times$ college-educated $_{t-1}$	0.486	0.223	0.474	0.228		
dependent variable:	unemplo	yed in t				
$high-paid_{t-1}$	-1.412	0.550	-1.698	0.560		
$\times$ college-educated $_{t-1}$	0.882	0.584	1.132	0.594		
$\times$ low social class <sub>t-1</sub>	0.533	0.588	0.474	0.579		
$low-paid_{t-1}$	-0.564	0.495	-0.497	0.484		
$\times$ college-educated $t-1$	-0.334	0.554	-0.226	0.547		
$\times$ low social class <sub>t-1</sub>	-0.385	0.482	-0.531	0.480		
unemployed $_{t-1}$	reference category					
$\times$ college-educated $_{t-1}$	0.065	0.356	0.117	0.358		
$\sigma_{\kappa_1}^2$	0.454	0.205	0.533	0.218		
$\sigma_{\kappa_1}^2$ $\sigma_{\kappa_2}^2$	0.827	0.584	0.944	0.578		
$\rho_{\kappa}$	_	_	0.737	0.260		
log likelihood	-579.379			-575.996		
observations	796		796			

Source: BHPS waves 8-18, own calculations. Coefficients displayed in bold are significant at least at the 10% level. Estimations include additional covariates as enlisted in Table 2 and vear dummies.

**Table:** Average Partial Effects

	Men with less than a college degree							
	RE Probit		CSM F	CSM RE Probit				
	APE	<i>p</i> -value	APE	<i>p</i> -value				
partial effect to obtain a high-paid employment in t								
$high-paid_{t-1}$	0.274	0.001	0.281	0.001				
$\times$ low social class <sub>t-1</sub>	0.217	0.002	0.231	0.001				
$low-paid_{t-1}$	0.111	0.177	0.163	0.048				
$\times$ low social class $_{t-1}$	0.033	0.580	0.097	0.134				
partial effect to obtain a low-paid employment in t								
$high-paid_{t-1}$	-0.067	0.287	-0.014	0.849				
$\times$ low social class $_{t-1}$	-0.050	0.341	0.000	0.957				
$low-paid_{t-1}$	0.003	0.907	-0.020	0.785				
$\times$ low social class $_{t-1}$	0.106	0.111	0.086	0.316				
partial effect to be unemployment in t								
$high-paid_{t-1}$	-0.206	0.009	-0.267	0.005				
$\times$ low social class $_{t-1}$	-0.167	0.009	-0.231	0.005				
$low-paid_{t-1}$	-0.114	0.111	-0.143	0.073				
$ imes$ low social class $_{t-1}$	-0.139	0.029	-0.183	0.011				
observations	143			143				

Source: BHPS waves 8-18, own calculations. APE=Average Partial Effect.





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# **Robustness:**

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1. Dropping observations with a relative wage change below < |10%|.

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- 1. Dropping observations with a relative wage change below < |10%|.
- 2. Gender specific threshold.
- 3. Including women into the regression.





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- 1. Low-wages increase the probability to switch into a high-paid employment.
- 2. Low-wages reduce the risk of future unemployment.
- **3.** Upward mobility is reduced when the job is associated with a low social class.
- **4.** Men with at least a college degree profit less strong from low-wages.
- **5.** Definition of low-wage threshold and sample composition have a strong influence on the transition probability between low-paid and high-paid.

# Thank you for your attention!